# New Modeling Methods for Multilevel Data 

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## Overview

- Two-level modeling for categorical data: weighted least squares estimation
- Three-level structural equation modeling: maximum-likelihood estimation
- Bayesian estimation for two and three level modeling
- Three-level modeling for categorical data: Bayesian estimation
- Cross classified structural equation models
- Two-level missing data imputation
- Two-level exploratory factor analysis
- Two-level multiple group modeling
- Plausible values for random effects


## Two-level modeling for categorical data: weighted least squares estimation

- Asparouhov, T. \& Muthén, B. (2007). Computationally efficient estimation of multilevel high-dimensional latent variable models. Proceedings of the 2007 JSM.
- Extends the single level methodology originated by Muthén (1978) Contributions to factor analysis of dichotomous variables.
- Mplus Version 5-2007
- Advantages over ML: unlimited number of latent variables, fast computation, useful chi-square test of fit
- Disadvantages over ML: models only continuous and categorical variables, MAR missing data not supported, can yield non-positive definite variance/covariance for large number of random effects, only random intercepts.

$$
\begin{gather*}
y_{p i j}=k \Leftrightarrow \tau_{p k}<y_{p i j}^{*}<\tau_{p k+1}  \tag{1}\\
y_{p i j}^{*}=y_{w p i j}+y_{b p j}  \tag{2}\\
y_{w i j}=\Lambda_{w} \eta_{w i j}+\varepsilon_{w i j}  \tag{3}\\
\eta_{w i j}=B_{w} \eta_{w i j}+\Gamma_{w} x_{w i j}+\xi_{w i j}  \tag{4}\\
y_{b j}=v_{b}+\Lambda_{b} \eta_{b j}+\varepsilon_{b j}  \tag{5}\\
\eta_{b j}=\alpha_{b}+B_{b} \eta_{b j}+\Gamma_{b} x_{b j}+\xi_{b j} \tag{6}
\end{gather*}
$$

## Two-level WLS: Estimation

Part 1. Estimate the unrestricted model

$$
\begin{gather*}
y_{p i j}=k \Leftrightarrow t_{p k}<y_{p i j}^{*}<t_{p k+1}  \tag{7}\\
y_{p i j}^{*}=y_{w p i j}+y_{b p j}  \tag{8}\\
y_{w i j}=\beta_{w} x_{w i j}+\varepsilon_{w i j}  \tag{9}\\
y_{b j}=\mu_{b}+\beta_{b} x_{b j}+\varepsilon_{b j} \tag{10}
\end{gather*}
$$

- ML Univariate estimation-1 dimension of numerical integration
- ML Bivariate estimation - 2 dimensions of numerical integration estimation only for 2 correlation parameters


## Two-level WLS: Estimation

Part 2. Estimate the structural model by minimizing

$$
\begin{equation*}
F=\left(s-s^{*}\right) W\left(s-s^{*}\right)^{T} \tag{11}
\end{equation*}
$$

$-s$ are the parameters of the unrestricted model
$-s^{*}$ are the implied quantities from the structural model

- $W$ is a weight matrix
- F is the base of the chi-square test of fit for the structural model v.s. the unrestricted model.


## Two-level WLS: Simulation Study

- 6 polytomous observed variables with 5 categories
- 100 clusters of size 10
- 2 within and 2 between factors in a two-level CFA each measured by 3 variables
- ML uses 8 dimensional integration - montecarlo integration


## Two-level WLS: Simulation Study Results

Table: Two-level factor analysis model with categorical variables.

| para- <br> meter | true <br> value | WLSM <br> bias | ML <br> bias | WLSM <br> coverage | ML <br> coverage | Efficiency <br> ratio |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\lambda_{w 2}$ | 1.0 | $3 \%$ | $2 \%$ | $97 \%$ | $100 \%$ | 1.14 |
| $\psi_{w 12}$ | 0.4 | $2 \%$ | $-14 \%$ | $97 \%$ | $89 \%$ | 0.89 |
| $\psi_{w 11}$ | 0.7 | $2 \%$ | $-23 \%$ | $94 \%$ | $75 \%$ | 0.71 |
| $\lambda_{b 2}$ | 1.0 | $5 \%$ | $4 \%$ | $96 \%$ | $94 \%$ | 0.96 |
| $\psi_{b 12}$ | 0.2 | $-1 \%$ | $-22 \%$ | $94 \%$ | $81 \%$ | 0.91 |
| $\psi_{b 11}$ | 0.4 | $1 \%$ | $-31 \%$ | $93 \%$ | $57 \%$ | 0.77 |
| $\tau_{11}$ | -0.3 | $-3 \%$ | $-6 \%$ | $96 \%$ | $87 \%$ | 1.17 |
| $\tau_{12}$ | 0.4 | $-1 \%$ | $-14 \%$ | $96 \%$ | $81 \%$ | 1.00 |
| $\tau_{13}$ | 1.2 | $0 \%$ | $-11 \%$ | $95 \%$ | $55 \%$ | 0.71 |
| $\tau_{14}$ | 1.8 | $0 \%$ | $-10 \%$ | $98 \%$ | $47 \%$ | 0.56 |
| $\theta_{b 1}$ | 0.2 | $-2 \%$ | $-55 \%$ | $97 \%$ | $32 \%$ | 0.66 |

## Three-level structural equation modeling: maximum-likelihood estimation

## 3-level SEM

- New feature in Mplus 7
- Three level multivariate data $Y_{p i j k}$

$$
\begin{equation*}
Y_{p i j k}=Y_{1 p i j k}+Y_{2 p j k}+Y_{3 p k} \tag{12}
\end{equation*}
$$

- 3 sets of structural equations - one on each level

$$
\begin{gather*}
Y_{1 i j k}=\Lambda_{1} \eta_{i j k}+\varepsilon_{i j k}  \tag{13}\\
\eta_{i j k}=B_{1} \eta_{i j k}+\Gamma_{1} x_{i j k}+\xi_{i j k} .  \tag{14}\\
Y_{2 j k}=\Lambda_{2} \eta_{j k}+\varepsilon_{j k}  \tag{15}\\
\eta_{j k}=B_{2} \eta_{j k}+\Gamma_{2} x_{j k}+\xi_{j k} .  \tag{16}\\
Y_{3 k}=v+\Lambda_{3} \eta_{k}+\varepsilon_{k}  \tag{17}\\
\eta_{k}=\alpha+B_{3} \eta_{k}+\Gamma_{3} x_{k}+\xi_{k} . \tag{18}
\end{gather*}
$$

## 3-level SEM: Estimation

- Double EM-algorithm where the latent variables are $Y_{2 j k}$ and $Y_{3 k}$
- E-step is based on $\left[Y_{3 k} \mid *\right]$ and $\left[Y_{2 j k} \mid Y_{3 k}, *\right]$
- M-step is a simple 3 group SEM maximization.
- Missing data: MAR support. Not a part of the EM-algorithm.
- Model accommodates easily variables defined on different levels: 7 types.
- Random coefficients for level 1 covariates and level 2 covariates

$$
\begin{gathered}
\Gamma_{1}=\gamma_{1}+\gamma_{1 j k}+\gamma_{1 k} \\
\Gamma_{2}=\gamma_{2}+\gamma_{2 k}
\end{gathered}
$$

where $\gamma_{1 j k}$ is a level 2 latent variable while $\gamma_{1 k}$ and $\gamma_{2 k}$ are level 3 latent variables.

- No numerical integration
- Fast computation - all runs less than 1 min.
- Robust standard errors using sandwich estimator.


## 3-level SEM Example 1: Path Analysis / 3-level regression

$$
\begin{gathered}
Y_{i j k}=Y_{1 i j k}+Y_{2 j k}+Y_{3 k} \\
Z_{i j k}=Z_{1 i j k}+Z_{2 j k}+Z_{3 k} \\
Y_{1 i j k}=\beta_{1} Z_{1 i j k}+\varepsilon_{i j k} \\
Y_{2 j k}=\beta_{2} Z_{2 j k}+\varepsilon_{j k} \\
Y_{3 k}=\alpha+\beta_{3} Z_{3 k}+\varepsilon_{k}
\end{gathered}
$$

- Simulation: 80 level 3 clusters, with 15 level 2 clusters each of size 10.
- Extension of BFSPE, bias reduction (observed v.s. latent predictor), and Mediation modeling to 3 level
- Marsh et al. (2009). Doubly-latent models of school contextual effects: Integrating multilevel and structural equation approaches to control measurement and sampling errors. MBR, 44, 764-802.
- Preacher et al. (2010). A general multilevel SEM framework for assessing multilevel mediation. Psychological Methods, 15, 209-233.


## 3-level SEM Example 1: Path Analysis Results

MODEL RESULTS


## 3-level SEM Example 2: Factor Analysis

- Simulation: 1 factor on each level with 5 indicators.
- 3 simulations with different number of level 3 clusters: 80,40 , 20.
- Chi-square test of fit between the restricted model factor analysis model and the unrestricted variance covariance.
- $\mathrm{DF}=15$, average chi-square value 15.156 / 15.571 / 17.707, rejection rates 5\% / 9\% / 19\%


## 3-level SEM Example 2: Factor Analysis Results, 80 level 3 clusters

MODEL RESULTS


## 3-level SEM Example 2: Factor Analysis Results, 40 level 3 clusters

MODEL RESULTS


Between CLUSTER2 Level

| EB2 BY |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Y2 | 1.000 | 1.0198 | 0.1996 | 0.2088 | 0.0398 | 0.950 | 1.000 |
| Intercepts |  |  |  |  |  |  |  |
| Y1 | 1.000 | 0.9838 | 0.1579 | 0.1535 | 0.0249 | 0.950 | 1.000 |
| Variances |  |  |  |  |  |  |  |
| EB2 | 0.600 | 0.5848 | 0.2036 | 0.2001 | 0.0413 | 0.920 | 0.950 |
| Residual Variances |  |  |  |  |  |  |  |
| Y1. | 0.300 | 0.2819 | 0.1076 | 0.0943 | 0.0118 | 0.810 | 0.900 |

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## 3-level SEM Example 2: Factor Analysis Results, 20 level 3 clusters

MODEL RESULTS
Population

ESTIMATES Average Std. Dev.
S. E.

Average
M. S. E. $95 \%$ \& sig Cover Coeff Within Level
E BY

## Variances

1.000
0.9948
0.0388
0.0387
$0.0015 \quad 0.9501 .000$

Residual Variances
Y1 1.200
1.1977
0.0382
0.0388
$0.0015 \quad 0.970 \quad 1.000$

Between CLUSTER1 Level


Between CLUSTER2 Level

| EB2 <br> Y2 | 1.000 | 1.0622 | 0.3626 | 0.3714 | 0.1340 | 0.900 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Intercepts <br> Y1 | 1.000 | 0.9687 | 0.2032 | 0.2147 | 0.0418 | 0.950 |
| Variances <br> EB2 | 0.600 | 0.5824 | 0.2963 | 0.2640 | 0.0872 | 0.830 |
| Residual Variances <br> Y1 | 0.300 | 0.2742 | 0.1450 | 0.1292 | 0.02150 .880 | 0.610 |

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## 3-level SEM Example 2: Factor Analysis with random coefficients

- Simulation: 5 observed indicator variables $Y_{p i j k}$ and one covariate $X_{i j k}$.

$$
\begin{gathered}
Y_{p i j k}=\mu_{p}+\lambda_{p} \eta_{i j k}+\varepsilon_{p i j k} \\
\eta_{i j k}=\eta_{0 i j k}+s_{j k} x_{i j k} \\
\eta_{0 i j k}=\eta_{1 i j k}+\eta_{2 j k}+\eta_{3 k} \\
s_{j k}=s_{2 j k}+s_{3 k}
\end{gathered}
$$

- 5 indicators variables measuring one factor regressed on a single predictor with random intercept and slope varying over level 2 and level 3
- 100 level 3 clusters, each with 20 level 2 clusters, each with 5 observations with 5 indicator variables


## 3-level SEM Example 2: Factor Analysis with random coefficients results

MODEL RESULTS


# Three-level structural equation modeling for categorical and continuous variable: Bayesian estimation 

## Bayesian Estimation

- The same 3-level structural model as with ML
- MCMC estimation: slopes / variances / latent variables / missing data / between level data
- The key is the likelihood posteriors $\left[Y_{3 k} \mid *\right]$ and $\left[Y_{2 j k} \mid Y_{3 k}, *\right]$ - the same as in the EM algorithm, followed by multiple group analysis
- Algorithm uses largest blocks possible for efficient mixing
- Advantages over ML: accommodates priors, not reliant on asymptotic theory, posterior for random effects and more accurate estimates, flexible modeling in structural part
- Easily extends modeling to categorical variables
- New feature in Mplus 7


## 3-level with categorical data: univariate regression

- $y_{i j k}$ categorical variable with 3 possible outcomes
- $x_{1 i j k}, x_{2 j k}, x_{3 k}$ covariates at level 1,2 and 3 .

$$
\begin{gathered}
y_{i j k}=l \Leftrightarrow \tau_{l}<y_{i j k}^{*}<\tau_{l+1} \\
y_{i j k}^{*}=\alpha_{j k}+\alpha_{k}+\beta_{1} x_{1 i j k}+\beta_{2} x_{2 j k}+\beta_{3} x_{3 k}+\varepsilon_{i j k} \\
\beta_{1}=\beta_{10}+\beta_{1 j k}+\beta_{1 k} \\
\beta_{2}=\beta_{20}+\beta_{2 k}
\end{gathered}
$$

- Variance of $\varepsilon_{i j k}$ is 1 as in probit regression
- ML would use 5 dimensional integration - very difficult
- Bayesian estimation $\leq 1$ min
- Default of weakly informative priors: IW(I,p+1), N(0,5). Key for small sample problems.


## 3-level with categorical data: univariate regression results

Between CLUSTER1 Level

| S | ON |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| X2 |  | 0.400 | 0.3938 | 0.0635 | 0.0550 | 0.0040 | 0.880 | 1.000 |
| Y | WITH |  |  |  |  |  |  |  |
| S |  | 1.000 | 0.9732 | 0.1201 | 0.1151 | 0.0150 | 0.880 | 1.000 |
| Residual Variances |  |  |  |  |  |  |  |  |
| Y |  | 1.500 | 1.4758 | 0.1676 | 0.1478 | 0.0284 | 0.910 | 1.000 |
| S |  | 1.600 | 1.5260 | 0.1850 | 0.1740 | 0.0393 | 0.920 | 1.000 |
| Between CLUSTER2 Level |  |  |  |  |  |  |  |  |
| S | ON |  |  |  |  |  |  |  |
| x3 |  | 0.550 | 0.5183 | 0.2106 | 0.2246 | 0.0449 | 0.950 | 0.620 |
| S2 | ON |  |  |  |  |  |  |  |
| X3 |  | 0.550 | 0.5646 | 0.2540 | 0.2135 | 0.0641 | 0.870 | 0.710 |
| Y | ON |  |  |  |  |  |  |  |
| X3 |  | 1.300 | 1.2793 | 0.1815 | 0.1795 | 0.0330 | 0.920 | 1.000 |
| S2 | WITH |  |  |  |  |  |  |  |
| S |  | 0.300 | 0.2596 | 0.2505 | 0.2412 | 0.0637 | 0.930 | 0.220 |
| Y |  | -0.200 | -0.2188 | 0.1844 | 0.1908 | 0.0340 | 0.920 | 0.180 |
| Y | WITH |  |  |  |  |  |  |  |
| $s$ |  | 0.700 | 0.6413 | 0.2068 | 0.2182 | 0.0458 | 0.950 | 0.940 |
| Intercepts |  |  |  |  |  |  |  |  |
| 5 |  | 1.700 | 1.6872 | 0.1582 | 0.1573 | 0.0249 | 0.940 | 1.000 |
| S2 |  | 0.300 | 0.2971 | 0.1497 | 0.1487 | 0.0222 | 0.930 | 0.540 |
| Thresholds |  |  |  |  |  |  |  |  |
| Y\$1 |  | -2.100 | -2.0968 | 0.1380 | 0.1152 | 0.0189 | 0.860 | 1.000 |
| Y\$2 |  | 2.100 | 2.0661 | 0.1600 | 0.1179 | 0.0265 | 0.820 | 1.000 |
| Residual Variances |  |  |  |  |  |  |  |  |
| Y |  | 1.300 | 1.2368 | 0.2444 | 0.2315 | 0.0631 | 0.910 | 1.000 |
| 5 |  | 2.200 | 2.0684 | 0.3702 | 0.3756 | 0.1530 | 0.930 | 1.000 |
| S2 |  | 2.000 | 1. 9354 | 0.3586 | 0.3428 | 0.1314 | 0.950 | 1.000 |

## 3-level SEM with categorical data: Factor model

- $y_{p i j k} 5$ categorical variable with 3 possible outcomes
- 1 factor on each level
- ML would use 13 dimensions of integration

$$
\begin{gathered}
y_{p i j k}=l \Leftrightarrow \tau_{p l}<y_{p i j k}^{*}<\tau_{p l+1} \\
y_{p i j k}^{*}=y_{p j k}+y_{p k}+\lambda_{1 p} f_{1 i j k}+\lambda_{2 p} f_{2 j k}+\lambda_{3 p} f_{3 k}+\varepsilon_{p i j k}
\end{gathered}
$$

- Variances of $\varepsilon_{p i j k}, f_{1 i j k}, f_{2 j k}, f_{3 k}$ are 1


## 3-level SEM with categorical data: Factor model results

| Population | ESTIMATES <br> Average | Std. Dev. | S. E. <br> Average | M. S. E. | $\begin{aligned} & 958 \\ & \text { Cover } \end{aligned}$ | \& sig Coeff |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Within Level |  |  |  |  |  |  |
| E1 BY |  |  |  |  |  |  |
| Y1 1.300 | 1.3130 | 0.0566 | 0.0496 | 0.0033 | 0.890 | 1.000 |
| Between CLUSTER1 Level |  |  |  |  |  |  |
| E2 BY |  |  |  |  |  |  |
| Y1 1.000 | 1.0094 | 0.0639 | 0.0564 | 0.0041 | 0.920 | 1.000 |
| Residual Variances |  |  |  |  |  |  |
| Y1 0.400 | 0.4042 | 0.0607 | 0.0573 | 0.0037 | 0.930 | 1.000 |
| Between CLuSter2 Level |  |  |  |  |  |  |
| E3 BY |  |  |  |  |  |  |
| $\begin{array}{ll}Y 1 & 0.600\end{array}$ | 0.6032 | 0.1056 | 0.1072 | 0.0111 | 0.940 | 1.000 |
| Thresholds |  |  |  |  |  |  |
| Y1\$1 -1.000 | -0.9994 | 0.1077 | 0.0866 | 0.0115 | 0.890 | 1.000 |
| Y1\$2 1.000 | 1.0180 | 0.1002 | 0.0876 | 0.0103 | 0.890 | 1.000 |
| Residual Variances |  |  |  |  |  |  |
| Y1 0.300 | 0.3139 | 0.0750 | 0.0772 | 0.0058 | 0.960 | 1.000 |

## Cross-classified structural equation modeling

## Cross-classified data

- $Y_{p i j k}$ is the $p$-th observation for person $i$ belonging to level 2 cluster $j$ and level 3 cluster $k$.
- Level 2 clusters are not nested within level 3 clusters
- Examples:
- To model income: individuals are nested within the same geographical location and are nested within occupation clusters
- Students are nested within schools and nested within neighborhoods
- Student performance scores are nested within students and within teachers
- Studies where observations are nested within persons and treatments/situations
- Studies where observations are nested within neighborhoods and interviewer
- Generalizability theory (Brennan, 2001; Cronbach, Rajaratnam, \& Gleser, 1963), Items are considered a random sample from a population of items.


## Cross-classified modeling

- Why do we need to model both sets of clustering?
- A model that ignores the dependence between observations in the same cluster yields misspecifications, underestimates SE, fails to discover the true predictor/explanotory effect stemming from the clusters
- Why do we need random effects and not fixed effects?
- Fixed effects possible: Tucker3
- Fixed effect for one set of clusters and random for the other: dummy variables for one set of clusters
- If the number of clusters is more than 10 - modeling with fixed effects may lead to too many parameters and less accurate model.


## Cross-classified modeling

Gonzalez, De Boeck, Tuerlinckx (2008) A Double-Structure Structural Equation Model for Three-Mode Data. Psychological Methods, 337-353.

| Variable | Tucker3 (a) | Tucker3 (b) | SEM-MTMM | 2sSEM |
| :---: | :---: | :---: | :---: | :---: |
| Interactions | [PSR] | [PSR] | [PS] [PR] | [PR] [SR] |
| Parameters |  |  |  |  |
| Persons | fixed | random | random | random |
| Situations | fixed | fixed | fixed | random |
| Responses | fixed | fixed | fixed | fixed |
| Dependence structure |  |  |  |  |
| Dependent | irrelevant | pairs (S, R) | pairs (S, R) | R |
| Independent | irrelevant | P | P | nonoverlapping pairs (P, S) |

SEM-MTMM $=$ structural equation model for multitrait-multimethod data; $2 \mathrm{sSEM}=$ double-structure structural equation model; $\mathbf{P}=$ persons; $\mathrm{S}=$ situations; $\mathrm{R}=$ responses.

## Cross-classified model

- Univariate model
- Both cluster levels contribute with random effects

$$
Y_{i j k}=Y_{1 i j k}+Y_{2 j}+Y_{3 k}
$$

- $Y_{2 j}$ and $Y_{3 k}$ are random effects for the two cluster levels


## Cross-classified model

- General SEM model

$$
Y_{p i j k}=Y_{1 p i j k}+Y_{2 p j}+Y_{3 p k}
$$

- 3 sets of structural equations - one on each level

$$
\begin{gathered}
Y_{1 i j k}=v+\Lambda_{1} \eta_{i j k}+\varepsilon_{i j k} \\
\eta_{i j k}=\alpha+B_{1} \eta_{i j k}+\Gamma_{1} x_{i j k}+\xi_{i j k} \\
Y_{2 j}=\Lambda_{2} \eta_{j}+\varepsilon_{j} \\
\eta_{j}=B_{2} \eta_{j}+\Gamma_{2} x_{j}+\xi_{j} \\
Y_{3 k}=\Lambda_{3} \eta_{k}+\varepsilon_{k} \\
\eta_{k}=B_{3} \eta_{k}+\Gamma_{3} x_{k}+\xi_{k}
\end{gathered}
$$

## Cross - classified model estimation

- The parameter $v$ and $\alpha$ can be used on any level
- Bayesian MCMC estimation
- The two sets of random effects $Y_{2 p j}, Y_{3 p k}$ are treated separately
- The two sets of random effects $\left[Y_{2 p j} \mid *, Y_{3 p k}\right]$ and $\left[Y_{3 p k} \mid *, Y_{2 p j}\right]$ are level 2 random effect
- Easily extends to categorical variables
- New feature in Mplus 7


## Cross-classified model, example 1: Factor model

- 1 factor at the individual level and 1 factor at each of the clustering levels, 5 indicator variables on the individual level

$$
y_{p i j k}=\mu_{p}+y_{p j}+y_{p k}+\lambda_{1 p} f_{1 i j k}+\lambda_{2 p} f_{2 j}+\lambda_{3 p} f_{3 k}+\varepsilon_{p i j k}
$$

- Variances of $f_{1 i j k}, f_{2 j k}, f_{3 k}$ are 1 , all loadings are estimated.
- 50 level 2 clusters, 50 level 3 clusters, 1 unit within each cluster intersection
- Total sample size 2500
- 1 unit on the within level would not work if the clusters were nested
- Estimation takes less than 1 min per replication


## Cross-classified model example 1: Factor model results

| Population | ESTIMATES <br> Average | Std. Dev. | S. E. <br> Average | M. S. E. | $95 \%$ Cover | $\begin{aligned} & \text { \& sig } \\ & \text { Coeff } \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Within Level |  |  |  |  |  |  |
| E BY |  |  |  |  |  |  |
| Y1 1.500 | 1.5018 | 0.0286 | 0.0324 | 0.0008 | 0.970 | 1.000 |
| Residual Variances |  |  |  |  |  |  |
| Y1 1.200 | 1.1999 | 0.0400 | 0.0437 | 0.0016 | 0.990 | 1.000 |
| Between CLUSTER1 Level |  |  |  |  |  |  |
| E2 BY |  |  |  |  |  |  |
| Y1 1.000 | 1.0000 | 0.2822 | 0.2865 | 0.0788 | 0.920 | 0.940 |
| Residual Variances |  |  |  |  |  |  |
| Y1 1.500 | 1.6146 | 0.3755 | 0.4754 | 0.1527 | 0.990 | 1.000 |
| Between CLUSTER2 Level |  |  |  |  |  |  |
| E3 BY |  |  |  |  |  |  |
| Y 10.800 | 0.8606 | 0.1538 | 0.1698 | 0.0271 | 0.950 | 1.000 |
| Intercepts |  |  |  |  |  |  |
| Y1 2.200 | 2.1561 | 0.2477 | 0.2900 | 0.0627 | 0.990 | 1.000 |
| Residual Variances |  |  |  |  |  |  |
| Y1 0.500 | 0.5517 | 0.1582 | 0.1647 | 0.0274 | 0.940 | 1.000 |

## Cross-classified model, example 2: Gonzalez's example

- Observations are nested within individual and several specific situations, the dependent variables are 4 binary outcomes
- 1 observation for each pair of clustering units

$$
y_{p j k}^{*}=y_{p j}+y_{p k}+\varepsilon_{p j k}
$$

- Variances of $\varepsilon_{p j k}$ is fixed to 1
- 100 level 2 clusters, 100 level 3 clusters
- Identical structural model for the two cluster random effects

$$
\begin{aligned}
y_{1 j} & =\beta_{1} y_{3 j}+\beta_{2} y_{4 j}+\varepsilon_{1 j} \\
y_{2 j} & =\beta_{3} y_{3 j}+\beta_{4} y_{4 j}+\varepsilon_{2 j} \\
y_{1 k} & =\beta_{1} y_{3 k}+\beta_{2} y_{4 k}+\varepsilon_{1 k} \\
y_{2 k} & =\beta_{3} y_{3 k}+\beta_{4} y_{4 k}+\varepsilon_{2 k}
\end{aligned}
$$

## Cross-classified model, example 2: Gonzalez's example



Figure 3. Graphical representation of the research questions. $a$, $b, c$, and $d$ are effect parameters.

Identical structural model is estimated for the two-sets of random effects.

## Cross-classified model, example 2: Gonzalez's example results



## Cross-classified model, example 2: Gonzalez's example results

| Y1 | ON |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| z1 |  | 0.300 | 0.3059 | 0.0891 | 0.0932 | 0.0079 | 0.940 | 0.910 |
| 22 |  | -0.300 | -0.2932 | 0.0870 | 0.0900 | 0.0075 | 0.970 | 0.940 |
| Y2 | ON |  |  |  |  |  |  |  |
| Z1 |  | 0.500 | 0.5046 | 0.0834 | 0.0877 | 0.0069 | 0.970 | 1.000 |
| z2 |  | -0.500 | -0.5045 | 0.1004 | 0.0939 | 0.0100 | 0.960 | 1.000 |
| Y1 | WITH |  |  |  |  |  |  |  |
| Y2 |  | 0.400 | 0.3929 | 0.0788 | 0.0814 | 0.0062 | 0.960 | 1.000 |
| 21 | WITH |  |  |  |  |  |  |  |
| z2 |  | 0.400 | 0.4056 | 0.0805 | 0.0816 | 0.0064 | 0.950 | 1.000 |
| Thresholds |  |  |  |  |  |  |  |  |
| Y1\$1 |  | 0.200 | 0.1998 | 0.1400 | 0.1461 | 0.0194 | 0.970 | 0.310 |
| Y2\$1 |  | -0.500 | -0.4932 | 0.1400 | 0.1453 | 0.0194 | 0.950 | 0.910 |
| 21\$1 |  | 0.200 | 0.2263 | 0.1517 | 0.1398 | 0.0235 | 0.930 | 0.350 |
| 22\$1 |  | -0.500 | -0.4675 | 0.1503 | 0.1285 | 0.0234 | 0.880 | 0.910 |
| Variances |  |  |  |  |  |  |  |  |
| 21 |  | 0.500 | 0.5208 | 0.0767 | 0.0811 | 0.0063 | 0.950 | 1.000 |
| 22 |  | 0.800 | 0.8133 | 0.1342 | 0.1255 | 0.0180 | 0.960 | 1.000 |
| Residual Variances |  |  |  |  |  |  |  |  |
| Y1 |  | 0.500 | 0.5053 | 0.0748 | 0.0802 | 0.0056 | 0.950 | 1.000 |
| Y2 |  | 0.800 | 0.7995 | 0.1277 | 0.1265 | 0.0161 | 0.960 | 1.000 |

## Two-level factor model with random loadings

$$
y_{p i j}=y_{p j}+\lambda_{p j} \eta_{i j}+\varepsilon_{p j k}
$$

Bayesian estimation, based on two explicit posterior distributions

$$
\begin{aligned}
& {\left[y_{p j}, \lambda_{p j} \mid *, \eta_{i j}\right]} \\
& {\left[\eta_{i j} \mid *, y_{p j}, \lambda_{p j}\right]}
\end{aligned}
$$

## Multiple imputations for two-level data

## Multiple imputations for two-level data

- Bayesian estimation of unrestricted H1 model

$$
y_{p i j}=y_{w p i j}+y_{b p j}
$$

- Unrestricted variance covariance matrix for $y_{w p i j}$
- Unrestricted means and variance covariance matrix for $y_{b p j}$
- Imputation of data based on the MCMC estimation
- Categorical data imputation based on $y_{p i j}^{*}$ and unrestricted correlation matrices
- PX methodology for estimating correlation matrices: Mplus 6.1


## Multiple imputations for two-level categorical data example

- Asparouhov and Muthén (2010) Multiple Imputation with Mplus
- $P$ categorical variables with 4 categories, $M$ clusters of size 30
- Two-level factor model with 1 factor at each level

$$
Y_{p i j}^{*}=\lambda_{w p} \eta_{w i j}+\lambda_{b p} \eta_{b j}+\varepsilon_{b p j}+\varepsilon_{w p i j}
$$

- Variance of $\varepsilon_{w p i j}$ is fixed to 1
- Generate MAR missing data using

$$
P\left(Y_{j} \text { is missing }\right)=\frac{1}{1+\operatorname{Exp}\left(-1.5+0.15 \sum_{k=26}^{30} Y_{k}\right)}
$$

- Analyze the data with WLSMV directly and WLSMV after imputation


## Multiple imputations for two-level categorical data results

Table: $M S E$ for the threshold parameters for two-level imputation with categorical variables.

| Number of <br> Clusters | Number of <br> Variables | Direct <br> WLSMV | H1 Imputed <br> WLSMV |
| :---: | :---: | :---: | :---: |
| 50 | 10 | 0.352 | 0.268 |
| 200 | 10 | 0.204 | 0.077 |
| 50 | 30 | 0.418 | 0.273 |
| 200 | 30 | 0.235 | 0.074 |

Imputation based results are more accurate. WLSMV supports MCAR, MARX but not MAR.

## Two-level exploratory factor analysis

## Two-level exploratory factor analysis

- Mplus 5
- Based on ML estimation for all continuous variables and WLS for continuous and categorical
- Estimate an unrotated solution

$$
Y_{i j}=\mu+\Lambda_{0 w} \eta_{0 w}+\Lambda_{0 b} \eta_{0 b}+\varepsilon_{i j}
$$

- $\Lambda_{0 w}$ and $\Lambda_{0 b}$ have fixed 0 above the diagonal
- $\eta_{0 w}$ and $\eta_{0 b}$ are standard normal
- Independent rotation of $\Lambda_{0 w}$ and $\Lambda_{0 b}$ based on Jennrich's gradient projection algorithm (GPA)
- GPA minimizes $f\left(\Lambda_{0 w} H_{w}^{-1}\right)$ and $f\left(\Lambda_{0 b} H_{b}^{-1}\right)$ over orthogonal rotations $H_{w}$ and $H_{b}$ where $f$ is the rotation criteria


## Two-level exploratory factor analysis

- Mplus uses as a default the Geomin rotation criteria

$$
f(\Lambda)=\sum_{i=1}^{p}\left(\prod_{j=1}^{m}\left(\lambda_{i j}^{2}+\varepsilon\right)\right)^{1 / m}
$$

- Chi-square test of fit for evaluating the model
- As a preliminary step in deciding the number of factors estimate within level EFA or between level EFA
- Within level EFA estimates unrestricted variance covariance on the between level and factor model on the within level only
- Between level EFA estimates estimates unrestricted variance covariance on the within level and a factor model on the between level only


## Two-level exploratory factor analysis example

## Exploratory Factor Analysis Of Aggression Items

Item Distributions for Cohort 3: Fall 1st Grade ( $\mathrm{n}=362$ males in 27 classrooms)

|  | Almost <br> Never <br> (scored as 1) | Rarely <br> (scored as 2) | Sometimes <br> (scored as 3) | Often <br> (scored as 4) | Very Often <br> (scored as 5) | Almost <br> Always <br> (scored as 6) |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Stubborn | 42.5 | 21.3 | 18.5 | 7.2 | 6.4 | 4.1 |
| Breaks Rules | 37.6 | 16.0 | 22.7 | 7.5 | 8.3 | 8.0 |
| Harms Others | 69.3 | 12.4 | 9.40 | 3.9 | 2.5 | 2.5 |
| Breaks Things | 79.8 | 6.60 | 5.20 | 3.9 | 3.6 | 0.8 |
| Yells at Others | 61.9 | 14.1 | 11.9 | 5.8 | 4.1 | 2.2 |
| Takes Others' <br> Property | 72.9 | 9.70 | 10.8 | 2.5 | 2.2 | 1.9 |
| Fights | 60.5 | 13.8 | 13.5 | 5.5 | 3.0 | 3.6 |
| Harms Property | 74.9 | 9.90 | 9.10 | 2.8 | 2.8 | 0.6 |
| Lies | 72.4 | 12.4 | 8.00 | 2.8 | 3.3 | 1.1 |
| Talks Back to <br> Adults | 79.6 | 9.70 | 7.80 | 1.4 | 0.8 | 1.4 |
| Teases Classmates | 55.0 | 14.4 | 17.7 | 7.2 | 4.4 | 1.4 |
| Fights With <br> Classmates | 67.4 | 12.4 | 10.2 | 5.0 | 3.3 | 1.7 |
| Loses Temper | 61.6 | 15.5 | 13.8 | 4.7 | 3.0 | 1.4 |

## Two-level exploratory factor analysis example

## Hypothesized Aggressiveness Factors

- Verbal aggression
- Yells at others
- Talks back to adults
- Loses temper
- Stubborn
- Property aggression
- Breaks things
- Harms property
- Takes others' property
- Harms others
- Person aggression
- Fights
- Fights with classmates
- Teases classmates


## Two-level exploratory factor analysis example



## Two-level exploratory factor analysis example

| Number of clusters |  |  | 27 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Average cluste | $r$ size | 13.407 |  |  |
| Estimated Intraclass Correlations for the Y Variables |  |  |  |  |  |
|  | Intraclass |  | Intraclass |  | Intraclass |
| Variable | - Correlation | Variable | Correlation | Variable | Correlation |
| U1 | 0.110 | U2 | 0.121 | U3 | 0.208 |
| U4 | 0.378 | U5 | 0.213 | U6 | 0.250 |
| U7 | 0.161 | U8 | 0.315 | U9 | 0.208 |
| U10 | 0.140 | U11 | 0.178 | U12 | 0.162 |
| U13 | 0.172 |  |  |  |  |

## Two-level exploratory factor analysis example

| Within-level | Between-level |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Factors | Factors | Df | Chi-Square | CEI | RMSEA |
| unrestricted | 1 | 65 | $66(p=0.43)$ | 1.000 | 0.007 |
| 1 | 1 | 130 | 670 | 0.991 | 0.107 |
| 2 | 1 | 118 | 430 | 0.995 | 0.084 |
| 3 | 1 | 107 | 258 | 0.997 | 0.062 |
| $4^{*}$ | 1 | 97 | 193 | 0.998 | 0.052 |

* $4^{\text {th }}$ factor has no significant loadings


## Two-level exploratory factor analysis example

|  | Within-Level Loadings |  |  | Between-Level Loadings |
| :--- | :---: | :---: | :---: | :---: |
|  | Property | Verbal | Person | General |
| Stubborn | 0.00 | $0.78^{*}$ | 0.01 | $\mathbf{0 . 6 5}^{*}$ |
| Breaks Rules | $0.31^{*}$ | $0.25^{*}$ | $0.32^{*}$ | $\mathbf{0 . 6 1 ^ { * }}$ |
| Harms Others and Property | $\mathbf{0 . 6 4 ^ { * }}$ | 0.12 | $0.25^{*}$ | $\mathbf{0 . 6 8 ^ { * }}$ |
| Breaks Things | $\mathbf{0 . 9 8 ^ { * }}$ | 0.08 | $-0.12^{*}$ | $\mathbf{0 . 9 8}^{*}$ |
| Yells At Others | 0.11 | $\mathbf{0 . 6 7}^{*}$ | 0.10 | $\mathbf{0 . 9 3 ^ { * }}$ |
| Takes Others' Property | $\mathbf{0 . 7 3 ^ { * }}$ | $-0.15^{*}$ | $0.31^{*}$ | $\mathbf{0 . 8 0 ^ { * }}$ |
| Fights | 0.10 | 0.03 | $\mathbf{0 . 8 6 ^ { * }}$ | $\mathbf{0 . 7 9 ^ { * }}$ |
| Harms Property | $\mathbf{0 . 8 1 ^ { * }}$ | 0.12 | 0.05 | $\mathbf{0 . 8 6 ^ { * }}$ |
| Lies | $\mathbf{0 . 6 0 ^ { * }}$ | $0.25^{*}$ | 0.10 | $\mathbf{0 . 8 6 ^ { * }}$ |
| Talks Back To Adults | 0.09 | $0.78^{*}$ | 0.05 | $\mathbf{0 . 8 1 ^ { * }}$ |
| Teases Classmates | 0.12 | $0.16^{*}$ | $\mathbf{0 . 5 9 *}$ | $\mathbf{0 . 8 3 ^ { * }}$ |
| Fights With Classmates | -0.02 | 0.13 | $\mathbf{0 . 8 8 ^ { * }}$ | $\mathbf{0 . 8 4 ^ { * }}$ |
| Loses Temper | -0.02 | $\mathbf{0 . 8 5 ^ { * }}$ | $\mathbf{0 . 0 5}$ | $\mathbf{0 . 8 7 ^ { * }}$ |

## Two-level multiple group modeling

## Two-level multiple group modeling

- Two types of grouping variable: between (private v.s. public schools), within (female v.s. males)
- Between grouping variable: straight forward, likelihood is sum of independent groups
- Within grouping variable: if you treat it as if the grouping variable is between the two groups within the cluster are independent - misspecification
- Modeling options: the between random effect is the same across groups OR more flexible the between random effects is different across groups
- Is teacher's effect on student performance the same for females and males?
- Are there two correlated random effects (teachers ability to relate to different gender) or is it one random effect (teachers ability) that may affect students performance differently?


## Two-level multiple group modeling

- Within level grouping variable - one random effect (assuming fixed variance of 1): twolevel mixture(with known class variable)

$$
Y_{i j g}=\mu_{g}+\beta_{g} Y_{j}+\varepsilon_{i j g}
$$

- Within level grouping variable - two random effects (assuming different variance across groups): twolevel mixture(with known class variable)

$$
Y_{i j g}=\mu_{g}+Y_{g j}+\varepsilon_{i j g}
$$

- Both use numerical integration currently
- Using 3 level modeling and treating the grouping variable as level 2 variable ( $Y_{g j}$ are correlated via $Y_{j}$ ) : no numerical integration

$$
Y_{i j g}=\mu_{g}+Y_{g j}+Y_{j}+\varepsilon_{i j g}
$$

- LRT to decide between the three models - if nested


## Plausible Values

## Plausible Values

- With Baysian estimation we can obtain posterior distribution for the random effects and plausible valeus as draws from the posterior
- Plausible values are more accurate than ML factor scores especially for small sample size
- Plausible values variability accounts for the variability in the parameter estimates
- Can be used for further analysis - comparing the random effects of different clusters, estimate difference, can be used for further modeling
- In 3 level modeling plausible values on level 2 accurately reproduce the correlation between level 2 random effects, not factor scores.

